Joint delay and angle of arrival estimation usingEM algorithm for multipath signals arriving atantenna array

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Abstract

We propose a novel approach to estimate the angles of arrival (AOA) and delays of the multipath signals from digitally modulated sources arriving at an antenna array. Our method uses a collection of estimates of a space-time vector channel The method avoids computationally expensive optimization search by using Expectation Maximization (EM) algorithm. The useful behavior of the proposed algorithm is verified by simulations.

Introduction

Source localization is an issue of interest in wireless communications In this system, mobiles emit signals that arrive at the base station via multiple paths with different angle of arrivals and delays which are in turn used for source $\frac{1}{2}$ bile) localization. This paper focuses on jointly estimating AOAs and relative time delays of multipath propagation signals emnating from a single source and received by antenna array. Parametric joint angle/delay estimation has been investigated itself in paper and one of the authors in and the authors in the paper of the second second in th rameters using MUSIC type algorithm, which though suboptimal, but involves only two dimentional search. In $\left[3\right]$, the authors used ESPRIT type algorithm to estimate jointly the parameters Wireless communications channel can be characterized by slow varying parameters and fast varying parameters The

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amplitudes and the relative phase of each path are fast varying and is subject to (Rayleigh) fading. The stationarity of the fading depends on the speed of the mobile If the mobile is moving fast enough then we can consider fading to be stationary within single time slot and are uncorrelated among slots In contrast AOAs and delays does not change over many slots even if mobile is at relatively high speed. Typically, a mobile moving at 30 m/s changes angular position only by 0.1° with respect to the base station 3 Kms away. Our method exploits the stationarity of angles and delays, as well as the independence of fading over many time slots We assume that AOAs and time delays do not change over multiple channel estimates because they are slow varying parameters. We also assume that antenna array response has known structure

2 Data Model

Consider a single usere case which transmit signal via multiple paths. The received base band signal at an M-element antenna array at time t in the n_{th} interval is $y^-(t) = |y^-(t)|$ $y^{**}(t)|_{\mathcal{T}}$ which can be written as

$$
y^{n}(t) = \sum_{l} S_{l}^{n} h^{n}(t - lT) + n^{n}(t)
$$
\n(1)

where T is symbol period and $n^{n}(t)$ is additive white Gaussian hoise. S_l^{n} is digital transmitted sequence. $n^-(t)$ is the channel The channel can be modeled \blacksquare as

$$
h^{n}(t) = \sum_{i}^{Q} a(\theta_{i}) \beta_{i}(n) g(t - \tau_{i})
$$
\n(2)

where Q are the total number of paths. Each path is parametrized by a DOA θ_i . times delay i_k dimensionemplexem path attenuation fading - in The array - and array measured - in The array response to the path from direction θ_i is $a(\theta_i)$ and $g(t)$ is known pulse shape (modulation waveform). After sampling at a rate P times the symbol rate. A block formulation is obtained by taking N snapshots, yielding

$$
Y^n = H^n S^n + N^n, n = 1, 2, \dots, S
$$
\n(3)

The first step is to estimate the channel impulse response from the user to the antenna array. This can be accomplished by using training sequence or using some blind method. Let $H^n = \Lambda(n)$ be the estimate of the true channel H^n . We have

$$
X(n) = \widehat{H^n} = H^n + V^n \tag{4}
$$

where V^+ is the estimation noise matrix. H^+ can be written as

$$
H^n = A(\theta) \operatorname{diag}(\beta(n)) G^T(\tau) \tag{5}
$$

where $A(\theta) = [a(\theta_1),...,a(\theta_0)], \rho(n) = [\rho_1(n),...,\rho_0(n)]$ and $G^-(\tau) = [g^-(\tau_1),...,g^-(\tau_0)].$ Applying vectorization operator to eq(4) and using eq(5) and writing the resulting equation in matrix form as

$$
x(n) = U(\theta, \tau)\beta(n) + v(n), n = 1, 2..., N
$$
 (6)

where

$$
U(\theta, \tau) = G(\tau) \circ A(\theta) = [u(\theta_1, \tau_1, \dots, u(\theta_Q, \tau_Q))]
$$
\n⁽⁷⁾

 $u(\theta, \tau)$ is called space time matrix. It is assumed that the structure of the space time matrix U is known. It is also assumed that path fadings are normally distributed with zero mean. Estimation noise V is white Gaussian. To be able to identify (θ, τ) using eq(6), we need $U(\theta, \tau)$ to be strictly tall and full column rank

-EM based JADE

Expectation Maximization (EM) $[5]$, is an iterative approach to Maximum Likelihood (ML) estimation. Each iteration is composed of two steps: an Expectation (E) step and a Maximization (M) step. The aim is to maximize the log likelihood $l(\psi; D)$, where ψ are the unknown parameters of the model and D are the data Suppose that this optimization problem would be simplied by the knowledge of the additional variable χ , known as missing or hidden data. The set \mathbb{D} , as \mathbb{D} is the complete data set in the same context D is same context. In the same context \mathbb{D} refered to as incomplete data set). Correspondingly, the loglikelihood function $l_c(\psi; D_c)$ is refered to as complete data likelihood. χ is chosen such that the function $l_c(\psi; D_c)$ would be easily maximized if χ were known. However, since χ is not obsevable, l_c is a random variable and cannot be maximized directly. Thus, the EM algorithm relies on integrating over the distribution of χ , with the auxiliary function $Q(\psi, \psi) = E_Y[\ell_c(\psi, D_c | D, \psi)]$, which is the expected value of the complete data likelihood, given the observed data D and the parameter φ computed at the previous neration. Intuitively, computing Q corresponds to filling the missing data using the knowledge of the observed data and previous parameters. The auxiliary function is deterministic and can be maximized. An EM algorithm iterates the following two steps for k until local or global maximum of the likelihood is found

Expectation: Compute

$$
Q(\psi; \psi^{(k)}) = E_{\chi}[l_c(\psi; D_c | D, \psi^{(k)}]
$$
\n(8)

Maximization: Update the parameters as

$$
\psi^{(k+1)} = arg_{max_{\psi}} Q(\psi; \psi^{(k)}), \qquad (9)
$$

In some cases, it is difficult to analytically maximize $Q(\psi;\psi^{++})$, as required by the M-step of the above algorithm, and we are only able to compute a new value ψ \cdots that produces an increase of Q at each iteration. In this case we have so called Generalized EM (GEM) algorithm. Given the model of $eq(6)$, we are now ready to denote data set Let z and the complete data set Let z and the complete data set Let z and the complete where x is estimated channel and - n hidden data is fading coecients and - n hidden data is fading coecients. We are now in position to start the derivation of the E-step of the EM algorithm.

the pdf of z as function of u can be written as product between the likelihood and priori distribution of - ields - i

$$
f(z;u) = f(x(n), \beta(n);u) = f(x(n)|\beta(n), u)f(\beta(n);u)
$$
 (10)

which can be further written as afterne converged fit it does not be at the control depend on the parameter u

$$
f(z;u) = K_1 \exp(\frac{1}{2\sigma^2}(x(n) - u\beta(n)))^H(x(n) - u\beta(n)))
$$
 (11)

having multiple channel estimation over slots in the slots in the slots in the slots in the slots is a set of t Estep given by

$$
Q(u; u^{(k)}) = E(\sum_{n=1}^{N} log f(z; u | x(n); u^{(k)}))
$$
\n(12)

which can be further expanded as

$$
Q(u; u^{(k)}) = \sum_{n=1}^{N} E((x(n)^H x(n) - x^H u \beta(n) - \beta(n)^H u^H x + \beta(n)^H u^H u \beta(n)) |x(n)|; u^{(k)})
$$
\n(13)

the above equation can be written as after omitting constant term

$$
Q(u; u^{(k)}) = \sum_{n=1}^{N} (-x(n)^H u \widehat{\beta(n)} - \widehat{\beta(n)}^H u^H x(n) + Trace(u^H u \beta(n) \widehat{\beta(n)}^H)
$$
 (14)

where $p(n) = E[p(n)|x(n);u^{\gamma-\gamma}]$, and

$$
\beta(n)\widehat{\beta(n)}^H = E(\beta(n\beta(n)^H | x(n); u^{(k)}) \tag{15}
$$

where E is expectation operator. The conditional mean will also be Gaussian. The conditional mean and second order moment is given by

$$
\widehat{\beta(n)} = R_{bb}u^H x(n) \tag{16}
$$

where

$$
R_{bb}^{-1} = u^H u + R_{\beta(n)\beta(n)}^{-1}
$$
 (17)

$$
R_{\beta(n)\beta(n)} = exp(-\beta(n)R_{\beta(n)\beta(n)}\beta(n))
$$
\n(18)

where covariance matrix $\mu(\mu|\mu)$ is assumed to be the summation

$$
\beta(n)\widehat{\beta(n)}^H = R_{bb} + \widehat{\beta(n)}\widehat{\beta(n)}^H
$$
\n(19)

The maximization step (updated parameter values) is given by equation the following equations to zero to find the stationary points i.e.,

$$
\frac{\partial Q}{\partial \tau}|_{\theta} = \frac{\partial Q}{\partial u} \left(\frac{\partial u}{\partial \tau}\right)|_{\theta} = 0 \tag{20}
$$

$$
\frac{\partial Q}{\partial \theta}|_{\tau} = \frac{\partial Q}{\partial u} \left(\frac{\partial u}{\partial \theta}\right)_{\tau} = 0
$$
\n(21)

where

$$
\frac{\partial Q}{\partial u} = \sum_{n=1}^{N} \left(2\widehat{\beta(n)}^{H} x(n) + 2\beta(n)\widehat{\beta(n)} u \right) \tag{22}
$$

These two equations can be solved iteratively to get updated value of the pa rameters to be estimated. The above fixed point equations are nonlinear and hence there could be local minima or saddle points. That solution should be chosen which gives maximum value of the likelihood function

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